

Probabilistic construction of continuous-time Markov chains P-F Rodriguez

Recall our definition of a continuous-time Markov chain (CTMC - really: pure jump process with no explosion)

Def A CTMC on a countable state space S is a collection $\{P_x\}_{x \in S}$ of prob. measures on Ω , the space of right-continuous S -valued trajectories with finitely many jumps in finite intervals (endowed with canonical coordinates $X_t: \Omega \rightarrow S$, $\omega \mapsto \omega(t)$, canonical σ -algebra \mathcal{F} , filtration $(\mathcal{F}_t)_{t \geq 0}$, and shifts $(\theta_t)_{t \geq 0}$) such that

$$(1) \quad P_x [X_0 = x] = 1$$

$$(2) \quad E_x [f(X_{t+h}) | \mathcal{F}_t] \stackrel{P_x\text{-a.s.}}{=} E_{X_t} [f(X_h)] \quad \forall x \in S, t, h \geq 0$$

$f: S \rightarrow \mathbb{R}$ bdd.

The goal of this note is to prove the following existence result:

Theorem Suppose that $c(\cdot)$ and $(p_{xy})_{x,y \in S}$ satisfy

$$(3) \quad c: S \rightarrow [0, \infty)$$

$$(4) \quad p_{xy} \geq 0 \quad \forall x, y \in S, \quad \sum_{y \in S} p_{xy} = 1, \quad x \in S \quad \text{and} \quad p_{xx} = \begin{cases} 1, & \text{if } c(x) = 0 \\ 0, & \text{if } c(x) > 0 \end{cases}$$

and that

$$(5) \quad \sum_{n \geq 0} c(Z_n)^{-1} = \infty, \quad P_x^{\mathbb{Z}}\text{-a.s.}, \quad \text{for } x \in S$$

where $\{Z_n; n \geq 0\}$ is the unique discrete-time Markov chain with transition probabilities $(p_{xy})_{x,y \in S}$, and $P_x^{\mathbb{Z}}$ stands for its canonical law started at x .

Then, there exists a CTMC (as defined above) with transition rates $c(x)$, $x \in S$ and transition probabilities $(p_{xy})_{x,y \in S}$, i.e., such that

$$(6) \quad P_x(T > t) = e^{-c(x)t}, \quad \text{for } x \in S, t \geq 0$$

(understood as: $T = \infty$, P_x -a.s. if $c(x) = 0$)

$$(7) \quad P_x[X_T = y] = p_{xy}, \quad \text{for } x, y \in S$$

where $T = \inf\{s \geq 0, X_s \neq X_0\}$ (with the convention $\inf \emptyset = \infty$) denotes the time of the first jump.

Remark (s)

- 1) Assumption (5) prevents explosions (i.e. accumulations of jumps in finite time). It is in fact equivalent to requiring
- (8) $\lim_{n \rightarrow \infty} T_n = \infty$, \tilde{P}_x -a.s. (with $\{T_n, n \geq 1\}$, \tilde{P}_x as, as defined below in (11), (12)).

For a proof of this equivalence, see for instance [Li] Thm 2.33, p.75. It is however immediately clear that (3), (4) alone are not sufficient in general (a notable exception being $c(x) = \lambda \in [0, \infty)$, all x , see 2) below) to guarantee existence. To see this, consider the following

Example: $S = \mathbb{N}$, $p_{k, k+1} = 1 \forall k \in \mathbb{N}$, $c(k) = (k+1)^2$

Then, defining $T_k \sim \text{Exp}(c(k))$, indep., on some (Ω, \mathcal{A}, P) and $N_t = \sup \{n \geq 1 : \sum_{k=1}^n T_k \leq t\}$, $t \geq 0$ ($\sup \emptyset = 0$)

we see that $\{N_t : t \geq 0\}$ escapes to infinity (in finite time):

$$P(N_t = \infty \text{ for some } t \in [0, \infty)) \\ = P\left(\sum_{k=1}^{\infty} T_k < \infty\right) = 1 \text{ since } \sum_{k=1}^{\infty} E[T_k] = \sum_{k=1}^{\infty} \frac{1}{(k+1)^2} < \infty,$$

$$\sum_{k=1}^{\infty} \text{var}(T_k) = \sum_{k=1}^{\infty} \frac{1}{(k+1)^4} < \infty \text{ (see e.g. [Du] Thm. 2.5.3.)}$$

- 2) As discussed in class, a drastic simplification occurs in the homogenous case ($c(x) = \lambda \in [0, \infty)$, all x). Then, the

spatial and temporal dynamics decouple, and one simply sets $P_x = (P_x^{\mathbb{Z}} \otimes P) \circ \tilde{X}_x^{-1}$, where

\uparrow $\tilde{X}_x = \tilde{Z}_{N_x}$ with \tilde{Z} the discrete time chain with trans prob $(p_{xy})_{x, y \in S}$, and canon. law $P_x^{\mathbb{Z}}$ (s.t. $P_x^{\mathbb{Z}}(Z_0 = x) = 1$) and $(N_t)_{t \geq 0}$ a Poisson process of rate λ under P).

(for $t > 0$)

- 3) As further discussed in class, the CTMC constructed in the above Thm is necessarily unique.

- 4) Condition (5) is easily seen to hold if for instance $\sup_{x \in S} c(x) < \infty$, or if \tilde{Z}_x is irreducible and recurrent. \square

Proof of the Theorem

We will assume that $c(x) > 0$ for all $x \in S$ (no absorbing states). The proof in the general case is similar, albeit notationally more cumbersome, and left as an exercise. We consider, on an auxiliary prob. space $(\tilde{\Omega}, \tilde{\mathcal{A}}, \tilde{P})$, a family $\{\tau_n(x), n \geq 0, x \in S\}$ of independent RV's satisfying

$$(9) \quad \tilde{P}(\tau_n(x) \geq t) = e^{-c(x)t}, \quad t \geq 0,$$

as well as, on $(\Omega^{\mathbb{Z}}, \mathcal{A}^{\mathbb{Z}}, P_x^{\mathbb{Z}})$, $x \in S$,

$$(10) \quad \{Z_n; n \geq 0\} \text{ the canonical discrete-time MC with trans prob. } (p_{xy})_{x,y \in S} \text{ and write}$$

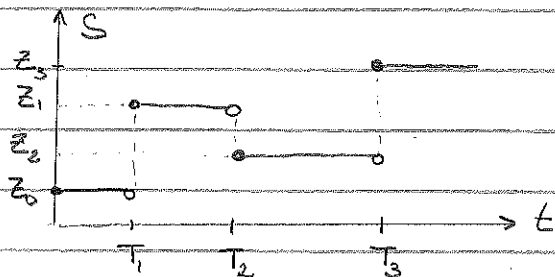
$$(11) \quad \tilde{P}_x \stackrel{\text{def.}}{=} P_x^{\mathbb{Z}} \otimes \tilde{P} \quad (\text{on } (\Omega^{\mathbb{Z}} \times \tilde{\Omega}, \mathcal{A}^{\mathbb{Z}} \times \tilde{\mathcal{A}})).$$

We then define the random variables (under \tilde{P}_x)

$$(12) \quad T_0 = 0, \quad T_1 = \tau_0(Z_0), \quad \dots, \quad T_n = \sum_{k=0}^{n-1} \tau_k(Z_k), \quad n \geq 1,$$

and set

$$(13) \quad \tilde{X}_t = \sum_{n \geq 0} Z_n \mathbb{1}_{\{T_n \leq t < T_{n+1}\}}, \quad \text{for } t \geq 0.$$



We will show that $\{P_x\}_{x \in S}$ with

$$(14) \quad P_x = \tilde{P}_x \circ (\tilde{X}_\cdot)^{-1} \quad (\text{the law of } (\tilde{X}_t)_{t \geq 0} \text{ under } \tilde{P}_x)$$

has the desired properties. First, by (5) and Remark 1), \tilde{X}_\cdot maps into Ω (of above (1)). Moreover,

$$P_x[X_0 = a] \stackrel{(14)}{=} \tilde{P}_x[\tilde{X}_0 = a] \stackrel{(13)}{=} P_x^{\mathbb{Z}}[Z_0 = a] = 1,$$

so (1) holds, and

$$P_x[T \geq t] \stackrel{(13)}{=} \tilde{P}_x[T_1 \geq t] \stackrel{(12)}{=} \tilde{P}_x[\tau_0(Z_0) \geq t] \stackrel{(9)}{=} e^{-c(x)t},$$

" $\underset{z \in P_x^{-1}(x)}{P_x^{\mathbb{Z}}}$

for $t \geq 0$, i.e. (6) holds, and (7) follows similarly. It thus remains to show $\{P_x\}_{x \in S}$ in (14) satisfy (2), which is the main task

Using Dynkin's π - λ -theorem, we see that (2) follows if we show

$$(15) \quad E_x [f_0(X_{s_0}) \cdots f_k(X_{s_k}) f(X_{s_k+h})] = E_x [f_0(X_{s_0}) \cdots f_k(X_{s_k}) E_{X_{s_k}} [f(X_{s_k+h})]]$$

$\forall x \in S, 0 \leq s_0 \leq \cdots \leq s_k, h > 0, f_i: S \rightarrow \mathbb{R}$ bdd

Let $N_t = \sup \{ n \geq 0, T_n \leq t \}$ (so that $X_t = Z_{N_t}$, see (13)),

then with $C = \{ N_{s_i} = n_i, i=0, \dots, k, N_{s_k+h} = n_k+n \}$

the left-hand side of (15) can be recast as

$$(16) \quad \text{LHS (15)} \stackrel{(14)}{=} \tilde{E}_x [f_0(Z_{N_{s_0}}) \cdots f_k(Z_{N_{s_k}}) f(Z_{N_{s_k+h}})] \\ = \sum_{\substack{0 \leq n_0 \leq n_1 \leq \cdots \leq n_k \\ n \geq 0}} \tilde{E}_x [\underbrace{\prod_{i=0}^k f_i(Z_{n_i}) f(Z_{n_k+n})}_{\stackrel{\text{def}}{=} F} 1_C]$$

The key is now to note that, conditionally on the Z 's, the dependence structure of the RV's N_{s_i} appearing in C is rather explicit, cf. (12): namely, cond on $Z_0, Z_1, \dots, Z_{n_k+n}$, the RV's $T_i(Z_i), 0 \leq i \leq n_k+n$ are indep., see (12) and (9) and $T_i(Z_i) \sim \text{Exp}(\lambda(Z_i)), 0 \leq i \leq n_k+n$. Hence, we can write

$$(17) \quad \tilde{E}_x [1_C | d(Z_0, \dots, Z_{n_k+n})] = h(Z_0, \dots, Z_{n_k+n})$$

where, for $x_0, \dots, x_{n_k+n} \in S$, $\stackrel{\text{def.}}{=} \int_{\mathbb{R}_+^{n_k+n}} d(x_0, \dots, x_{n_k+n})(t)$

$$(18) \quad h(x_0, \dots, x_{n_k+n}) = \int_A \left(\prod_{i=0}^{n_k+n} \lambda(x_i) e^{-\lambda(x_i)t_i} dt_i \right)$$

$$(19) \quad \text{with } A \stackrel{\text{def.}}{=} \left\{ (t_0, \dots, t_{n_k+n}) \in \mathbb{R}_+^{n_k+n} : t_0 + \dots + t_{k-1} \leq s_k \leq t_0 + \dots + t_{n_i}, 0 \leq i \leq k \right. \\ \left. \text{and } t_0 + \dots + t_{n_k+n-1} \leq s_k+h < t_0 + \dots + t_{n_k+n} \right\}$$

We now integrate over t_{n_k} in (18) using the memoryless property of the exponential, which can be stated as

$$(20) \quad \int_u^\infty \lambda e^{-\lambda t} \varphi(t-u) dt = e^{-\lambda u} \int_0^\infty \lambda e^{-\lambda t} \varphi(t) dt, \text{ for } u \geq 0, \lambda > 0 \\ \varphi \text{ bdd, meas.}$$

so that, with $t = t_{n_k}, \lambda = \lambda(x_{n_k}), u = s_k - (t_0 + \dots + t_{n_k-1})$

(which satisfies $u \geq 0$ on A), and noting moreover that the second line in (19) can be recast as

$$h - (t_{n_{k+1}} + \dots + t_{n_{k+n}}) \leq t_{n_k} - u \leq h - (t_{n_{k+1}} + \dots + t_{n_{k+n-1}})$$

we obtain:

$$(21) \quad h(x_0, \dots, x_{n_{k+n}}) = h_1(x_0, \dots, x_{n_k}) h_2(x_{n_k}, \dots, x_{n_{k+n}})$$

where

$$h_1(x_0, \dots, x_{n_k}) = \mu_{x_0, \dots, x_{n_k}} \left(\left\{ (t_0, \dots, t_{n_k}) : \begin{array}{l} t_0 + \dots + t_{i-1} \leq s_i \\ < t_0 + \dots + t_i, \quad 0 \leq i \leq k \end{array} \right\} \right)$$

$$h_2(x_{n_k}, \dots, x_{n_{k+n}}) = \mu_{x_{n_k}, \dots, x_{n_{k+n}}} \left(\left\{ (t_0, \dots, t_n) : \sum_{k=0}^{n-1} t_k \leq h < \sum_{k=0}^n t_k \right\} \right).$$

Substituting (21) into (17) and returning to (16) yields

$$\tilde{E}_x [F1_C] = E_x^Z [F h_1(Z_0, \dots, Z_{n_k}) h_2(Z_{n_k}, \dots, Z_{n_{k+n}})]$$

$$\stackrel{\text{!}}{\rightarrow} f_0(Z_{n_0}) \dots f_k(Z_{n_k}) f(Z_{n_{k+n}})$$

(only the randomness of Z remains!)

$$(22) \quad \begin{array}{l} \text{simple} \\ \text{Markov for } Z \\ \text{at time } n_k \end{array} \rightarrow E_x^Z \left[\prod_{i=0}^k f_i(Z_{n_i}) \underbrace{E_{Z_{n_k}}^Z [f(Z_{n_k}) h_2(Z_0, \dots, Z_{n_k})]}_{\stackrel{\text{def}}{=} G_n(Z_{n_k})} \right]$$

By a similar calculation as in (17), (18), but simpler, we find that, for $\hat{C} = \{N_{s_0} = n_0, \dots, N_{s_k} = n_k\}$,

$$(23) \quad E_x^Z [1_{\hat{C}} | \sigma(Z_0, \dots, Z_{n_k})] = h_1(Z_0, \dots, Z_{n_k})$$

so that, going back to (22), we obtain that

$$\tilde{E}_x [F1_C] = \tilde{E}_x \left[\prod_{i=0}^k f_i(Z_{n_i}) G_n(Z_{n_k}) 1_{\hat{C}} \right]$$

and, inserting this into (16), we finally get that

$$(24) \quad \text{LHS (15)} = \sum_{n \geq 0} \tilde{E}_x \left[\prod_{i=1}^k f_i(Z_{N_{s_i}}) G_n(Z_{N_{s_k}}) \right]$$

Yet another calculation as in (23) (but even simpler!) identifies

$$h_2(Z_0, \dots, Z_n) = \tilde{E}_x [1_{\{N_n = n\}} | \sigma(Z_0, \dots, Z_n)], \text{ which implies}$$

(cf. the def of $G_n(\cdot)$ in (22)) that

$$(25) \quad G_n(y) = \tilde{E}_y [f(Z_{N_n}) 1_{\{N_n = n\}}], \text{ for } n \geq 0, \text{ yes.}$$

Substituting (25) into (24) and summing over n , we see that

$$\text{LHS (15)} = \tilde{E}_x \left[\prod_{i=0}^k f_i(Z_{N_{s_i}}) \tilde{E}_{Z_{N_{s_k}}} [f(Z_{N_k})] \right]$$

$$\stackrel{\text{below (15)}}{=} \tilde{E}_x \left[\prod_{i=0}^k f_i(\tilde{X}_{s_i}) \tilde{E}_{\tilde{X}_{s_k}} [f(\tilde{X}_k)] \right],$$

which is the RHS of (15), on account of (14)! \square

References:

[Li] T.M. Liggett, Continuous-time Markov Processes - An Introduction. GSM, Vol 113, AMS, 2010.

[Du] R. Durrett, Probability, Theory and Examples. Cambridge University Press, 2010.