

Existence

It is enough to prove it exists for every interval  $[0, \alpha]$ ,  $\alpha > 0$ . This because the uniqueness tells us that the decompositions must agree where they should - for differing  $\alpha_1, \alpha_2$  - and we can therefore define the decomposition for any particular  $t \in \mathbb{R}^+$  simply by referring to the decomposition for any  $\alpha > t$ .

So fix  $\alpha$ . For  $0 \leq t \leq \alpha$ ,  $(M_t^t(X_\alpha))$  is a martingale with respect to the (appropriate) filtration. It admits a right continuous modification. Choose one. Consider

$$Y_t = X_t - M_t(X_\alpha) \quad 0 \leq t \leq \alpha$$

where  $(M_t(X_\alpha))$  denotes the chosen modification.

Let  $\pi^n = \{t_i^n\}$  denote a sequence of partitions of  $[0, \alpha]$ . Using the discrete analysis from the beginning, for each  $n \in \mathbb{N}$  we have the (Doob)

decompositions,

$$Y_{t_i^n}^{t_i^n} = U_{t_i^n}^{t_i^n} + A_{t_i^n}^{t_i^n}$$

where  $(A_{t_i^n}^{t_i^n})$  is a predictable increasing process. Now

$$Y_\alpha = 0 \quad \Delta_0 U_{t_i^n}^{t_i^n} = -A_{t_i^n}^{t_i^n} \quad (t_i^n \text{ is just } \alpha)$$

as a consequence,

$$Y_{t_i^n}^{t_i^n} = A_{t_i^n}^{t_i^n} - M_{t_i^n}^{t_i^n}(A_\alpha)$$

We can take  $\pi^n$  to be an increasing sequence of partitions with  $\|\pi^n\| \rightarrow 0$  as  $n \rightarrow \infty$ . For instance,  $\pi^n = \{j/2^n\}, 0 \leq j \leq 2^n$ .

What we do now is to show the sequence  $\{A_n^\alpha : n \in \mathbb{N}\}$  is uniformly integrable. (H) let  $\lambda > 0$ , and

\*\*  $T_n^\lambda = \alpha \vee \min\{t_n^{l_1} : A_n^{t_n} > \lambda \text{ for some } i\}$ .

Then for a particular index,  $i$ , we  $\{T_n^\lambda(\omega) \leq t_n^{l_1}\} \Leftrightarrow A_n^{t_n}(\omega) > \lambda$  for some  $j \leq i-1$  and therefore  $A_n$  is increasing. On the other hand, if we  $\{A_n^{t_n}(\omega) > \lambda\}$  then  $T_n^\lambda(\omega) \leq t_n^{l_1}$ . So  $\{T_n^\lambda(\omega) \leq t_n^{l_1}\} = \{A_n^{t_n}(\omega) > \lambda\}$  and since  $A_n$  is predictable  $\{T_n^\lambda(\omega) \leq t_n^{l_1}\} \in \mathcal{F}_{t_n^{l_1}}$ .

So  $T_n^\lambda$  is a stopping time, R-a.s. less than  $\alpha$ . [Whosps!]  
Note also  $\{T_n^\lambda < \alpha\} = \{A_n^{t_n} > \lambda\} \equiv \{A_n^\alpha > \lambda\}$  (as  $t_n^{l_1} = \alpha$ )

← This bit goes here.

Recall the definition of the  $\sigma$ -field  $\mathcal{F}_{T_n}^\lambda$ ;  $\{E\mathcal{F} : E\mathcal{F} \leq T_n^\lambda\}$   $E\mathcal{F}_t$ , for all  $t$ ; it is not difficult to prove that the conditional expectation  $M_{T_n}^\lambda$  may be written

$$M_{T_n}^\lambda = \sum_{T_n^\lambda = t_n^i} M_{t_n^i}^\lambda(\cdot) \mathbb{I}_{\{T_n^\lambda = t_n^i\}}$$

So, on  $\{T_n^\lambda = t_n^i\}$  the expectation  $M_{T_n}^\lambda$  agrees with  $M_{t_n^i}^\lambda$ . Now, recall that earlier we wrote

$$V_{t_n^i}^\lambda = A_n^{t_n^i} - M_{t_n^i}^\lambda(A_n^\alpha)$$

(F) Actually, what is used is weak relative compactness. I've got a direct proof, somewhere. Can't remember how it goes. \*\*  $T_n^\lambda$  is the "last time"  $A_n$  is less than  $\lambda$  and is  $\alpha$  if it's never greater.

It follows that

$$Y_{T_\lambda} = \sum_{t_i \in T_\lambda} Y_{t_i} I_{\{T_\lambda = t_i\}} = \sum_{t_i \in T_\lambda} A_{t_i}^n I_{\{T_\lambda = t_i\}} - \sum_{t_i \in T_\lambda} M_{t_i}^n(A_\alpha) I_{\{T_\lambda = t_i\}}$$

so that

$$Y_{T_\lambda} = A_{T_\lambda}^n - M_{T_\lambda}^n(A_\alpha)$$

$$\leq \lambda - M_{T_\lambda}^n(A_\alpha) \quad (\#)$$

So, as  $\{T_\lambda < \alpha\} = \{A_\alpha^{T_\lambda} > \lambda\}$ ,

$$\int_{\{T_\lambda < \alpha\}} Y_{T_\lambda} dP \leq \lambda P\{T_\lambda < \alpha\} - \int_{\{T_\lambda < \alpha\}} M_{T_\lambda}^n(A_\alpha) dP$$

once  $\{T_\lambda < \alpha\} \in \mathcal{F}_{T_\lambda}^n$  we can remove  $M_{T_\lambda}^n$  and rewrite and rearrange to get:

$$\int_{\{A_\alpha^{T_\lambda} > \lambda\}} A_\alpha dP = \lambda P\{T_\lambda < \alpha\} - \int_{\{T_\lambda < \alpha\}} Y_{T_\lambda} dP$$

The term on the left is what we're interested in. Uniform integrability is, roughly speaking, that the integrals of the family  $(A_\alpha^n)_{n \in \mathbb{N}}$  over a 'small' set should be (uniformly) small. If we can show that the left side above is small-uniformly in  $n \in \mathbb{N}$ , while  $\lambda$  gets large then this is enough.....

(+) Remember,  $T_\lambda$  is the last time  $A_\alpha^n$  is less than  $\lambda$

To do this we need some estimate of the right side of the last equation. Well, recall

$$Y_{T_n^\lambda} = A_n^{T_n^\lambda} - M_{T_n^\lambda}^{T_n^\lambda}(A_n^\alpha)$$

Trick: Replace  $\lambda$  with  $\frac{\lambda}{2}$  ( $\lambda$  was simply a positive value) and integrate over the set  $\{T_n^{\lambda/2} < \alpha\}$ , to get

$$-\int_{\{T_n^{\lambda/2} < \alpha\}} Y_{T_n^{\lambda/2}}^{T_n^{\lambda/2}} dP = \int_{\{T_n^{\lambda/2} < \alpha\}} \underbrace{(A_n^\alpha - A_n^{T_n^{\lambda/2}})}_{\text{this is non-negative}} dP$$

Now  $\{T_n^{\lambda/2} < \alpha\} \supseteq \{T_n^\lambda < \alpha\}$  (because, if  $T_n^\lambda = t_n^\lambda$  - the last time  $A_n^\bullet(\omega)$  is less than  $\lambda$  - then the last time it was less than  $\lambda/2$  occurred no later than this). So,

$$-\int_{\{T_n^{\lambda/2} < \alpha\}} Y_{T_n^{\lambda/2}}^{T_n^{\lambda/2}} dP \geq \int_{\{T_n^\lambda < \alpha\}} (A_n^\alpha - A_n^{T_n^{\lambda/2}}) dP$$

$$= \int_{\{A_n^\alpha > \lambda\}} A_n^\alpha dP - \int_{\{T_n^\lambda < \alpha\}} A_n^{T_n^{\lambda/2}} dP$$

The first of the right hand terms is greater than  $\lambda P\{T_n^\lambda < \alpha\}$  while the second is smaller than  $\frac{\lambda}{2} P\{T_n^\lambda < \alpha\}$  so the difference on the right side above is greater than

$$\lambda P\{T_n^\lambda < \alpha\} - \frac{\lambda}{2} P\{T_n^\lambda < \alpha\} = \frac{\lambda}{2} P\{T_n^\lambda < \alpha\}.$$

So we now have

$$-\int_{\{T_n^{1/2} < \alpha\}} Y_n^{T_n^{1/2}} dP \geq \frac{1}{2} \lambda P\{T_n < \alpha\}$$

This gives us an estimate for  $\lambda P\{T_n < \alpha\}$ !

$$\lambda P\{T_n < \alpha\} \leq -2 \int_{\{T_n^{1/2} < \alpha\}} Y_n^{T_n^{1/2}} dP,$$

giving the estimate

$$\int_{\{A_n^\alpha > \lambda\}} A_n^\alpha dP \leq -2 \int_{\{T_n^{1/2} < \alpha\}} Y_n^{T_n^{1/2}} dP - \int_{\{T_n < \alpha\}} Y_n^{T_n} dP$$

By hypothesis  $\{X_T : T \text{ a.s.t.m.s.} \leq \alpha\}$  is uniformly integrable.

Since  $Y_T = X_T - M_T(X)$  then  $\{Y_T : T \text{ a.s.t.m.s.} \leq \alpha\}$

is also uniformly integrable, so the sequences  $(Y_n^{T_n})$  and  $(Y_n^{T_n^{1/2}})$  are uniformly integrable. Finally,

$$P\{T_n < \alpha\} = P\{A_n^\alpha > \lambda\} \leq \frac{E(A_n^\alpha)}{\lambda} = \frac{E(Y_n)}{\lambda}$$

(the last bit is because  $\lambda I_{\{A_n^\alpha > \lambda\}} \leq A_n^\alpha I_{\{A_n^\alpha > \lambda\}}$  (integrate) and  $Y_n^{T_n} = A_n^{T_n} - M_n^{T_n}(A_n^\alpha)$  so  $Y_n = A_n - M_n(A_n^\alpha)$ ,  $A_n \equiv 0$ , take expectations)

This last inequality shows that  $\mathbb{P}\{T_n < \lambda\} \rightarrow 0$  as  $\lambda \rightarrow \infty$  uniformly in  $n$ , this estimates the term,

$$\int A_n^\alpha d\mathbb{P}, \quad \int \{A_n^\alpha > \lambda\}$$

showing it to be small (uniformly in  $n \in \mathbb{N}$ ) for large values of  $\lambda$ . As we remarked previously this proves that  $(A_n^\alpha)$  is uniformly integrable. Now (ii) uniform integrability entails weak relative compactness which means that there is an integrable random variable  $A^\alpha$  and a subsequence  $(A_{n_i}^\alpha)$  converging weakly in  $L^1$  to  $A^\alpha$ . This means

$$\lim_{j \rightarrow \infty} \mathbb{E}(A_{n_j}^\alpha B) = \mathbb{E}(A^\alpha B)$$

for  $B \in \mathcal{L}^\infty(\Omega, \mathcal{F}, \mathbb{P})$ . Now we define

$$A_t^\alpha = Y_t^\alpha + M_t^\alpha(A^\alpha) \quad 0 \leq t \leq \alpha.$$

Now let  $\pi = \bigcup_{n=1}^{\infty} \pi_n$ , a dense subset of  $[0, \alpha]$ .

Recalling that

$$Y_{t_n}^\alpha = A_{t_n}^\alpha - M_{t_n}^\alpha(A^\alpha)$$

for every  $t_n$  (and every  $n \in \mathbb{N}$ ) we note that for any

(i) See Dunford + Schwartz: Linear Operators Vol I, but don't linger, the point is what follows next.

$$\lim_{n_j} E(A^j B) = E(A^t B) \text{ for } B \in \mathcal{L}(\Omega, \mathcal{F}, \mathbb{P})$$

This is obvious but the notation obscures the fact:

The partitions  $\pi^n$  get larger and larger as  $n$  gets big. So a  $t \in \pi^n$  occurs in  $\pi^{n+1}, \pi^{n+2}, \dots$ , etc, but it will have different 'names' for each  $n \in \mathbb{N}$ , it will be  $t_n^i$  in  $\pi^{n+1}$ , but it will be  $t_n^i$  in  $\pi^{n+1}$ ,  $t_n^i$  in  $\pi^{n+2}$ , and so on. So for this bit of the argument let's say that  $t = t_n^{i(n)}$  in  $\pi^n$ . By definition,

$$A_t = Y_t + M_t(A^\alpha)$$

Take the right side

$$= A_n^{t_n^{i(n)}} - M_n^{t_n^{i(n)}}(A^\alpha) + M_t(A^\alpha)$$

of course  $t_n^{i(n)}$  is always  $t$ , but, a big but,  $A_n^{t_n^{i(n)}}$  changes with  $n$ . Now for  $B \in \mathcal{L}(\Omega, \mathcal{F}, \mathbb{P})$ , we have, in particular

$$E(A^t B) = E(A_n^{t_n^{i(n)}} B) - E(M_n^{t_n^{i(n)}}(A^\alpha) B)$$

$$E(M_t(A^\alpha) B)$$

lets drop the difference between  $t$  and  $t_n^{i(n)}$ , we get,

$$E(A^t B) = E(A_n^{t_n^{i(n)}} B) - E(M_n^{t_n^{i(n)}}(A^\alpha) B) - E(M_t(A^\alpha) B)$$



Of course

$$M_t(B) \in \mathcal{L}(\Omega, \mathcal{F}, \mathbb{P}) \quad \forall t$$

$$\lim_{n_j} E((A_{n_j}^\alpha - A^\alpha) M_t(B)) = 0$$

therefore  $E(A_t^t B) = \lim_{n_j} E(A_{n_j}^t B)$ . So now let  $0 \leq \Delta \leq t \leq \alpha$  and  $B_{n_j}$  as before but non negative

$$E((A_t - A_0) B) = \lim_{n_j} E((A_{n_j}^t - A_{n_j}^0) B) \geq 0 \quad (\dagger)$$

Setting  $B = I_{\{A_0 > A_t\}}$  then  $E((A_t - A_0) I_{\{A_0 > A_t\}}) \geq 0$

but, clearly  $E((A_t - A_0) I_{\{A_0 > A_t\}}) \leq 0$ . So

$E((A_t - A_0) I_{\{A_0 > A_t\}}) = 0$  and  $(A_t - A_0) I_{\{A_0 > A_t\}} = 0$

$\mathbb{P}$  a.e. If  $\mathbb{P}\{A_0 > A_t\} > 0$  then  $E((A_t - A_0) I_{\{A_0 > A_t\}}) < 0$

$\Rightarrow \mathbb{P}\{A_0 > A_t\} = 0$  and  $\mathbb{P}\{A_t \geq A_0\} = 1$ . So  $A_t \geq A_0$   $\mathbb{P}$  a.s. Now,  $\Pi$  is countable, by throwing away a large enough null set we have,  $\mathbb{P}$  a.s.

$$\exists t \mapsto A_t(\omega)$$

is monotone increasing. But  $(Y_t)$  and  $(M_t(A^\alpha))$  are  $\mathbb{P}$  a.s. right cts and so, therefore,  $(A_t)$  is  $\mathbb{P}$  a.s. It follows by right continuity that  $(A_t)$  is  $\mathbb{P}$  a.s. By definition

$$A_0 = Y_0 + M_0(A^\alpha)$$

$$= X_0 - M_0(X^\alpha) + M_0(A^\alpha)$$

(†) This is enough to show  $A_t \geq A_0$   $\mathbb{P}$  a.s. because it says  $A_t - A_0 \geq 0$  in  $L^1$ , but read on.



But remember " $Y_{t_i}^n = A_{t_i}^n - M_{t_i}^n(A^n)$ " so that

$Y_0 = A_0^n - M_0^n(A^n)$  for every  $n$ , and  $A_0^n \equiv 0$  in  $\mathbb{R}^d$ .

So  $Y_0 = M_0(A^n)$  in  $\mathbb{R}^d$ , and we see that  $A_0 = 0$  in  $\mathbb{R}^d$ .

Is  $(A_t)$  natural? Consider, a rt do modification of  $(M_t(B))$  for  $B \in \mathcal{L}_\infty(\Omega, \mathcal{F}, \mathbb{P})$ ; using a telescoping argument and predictability of  $(A_{t_i}^n)$  we get;

$$E(A_{t_n}^n M_{t_n}^n(B)) = E\left(\sum_{i=1}^n M_{t_{i-1}}^n(B)(A_{t_i}^n - A_{t_{i-1}}^n)\right)$$

$$= E\left(\sum_{i=1}^n M_{t_{i-1}}^n(B)(Y_{t_i}^n - Y_{t_{i-1}}^n)\right)$$

because  $A_{t_i}^n = Y_{t_i}^n + M_{t_i}^n(A^n)$  and the martingale part will vanish inside the expectation.

$$= E\left(\sum_{i=1}^n M_{t_{i-1}}^n(B)(A_{t_i}^n - A_{t_{i-1}}^n)\right)$$

because  $A_{t_i}^n = Y_{t_i}^n + M_{t_i}^n(A^n)$  and, once again, when we have written  $Y$  in terms of  $A$  and  $M$  the martingale part,  $(M_{t_i}^n(A^n))$ , will vanish inside the expectation. Taking a limit gives

$$E(A_{t_n}^n M_{t_n}^n(B)) = E \int_0^{t_n} M_{s-}^n(B) dA_s.$$

(Proposition 2).

Finally:

$$A_t = Y_t + M_t(A^n) = X_t - M_t(X) + M_t(A^n)$$

so,

$$X_t = M_t(X - A^n) + A_t$$

on  $[0, \alpha]$ .